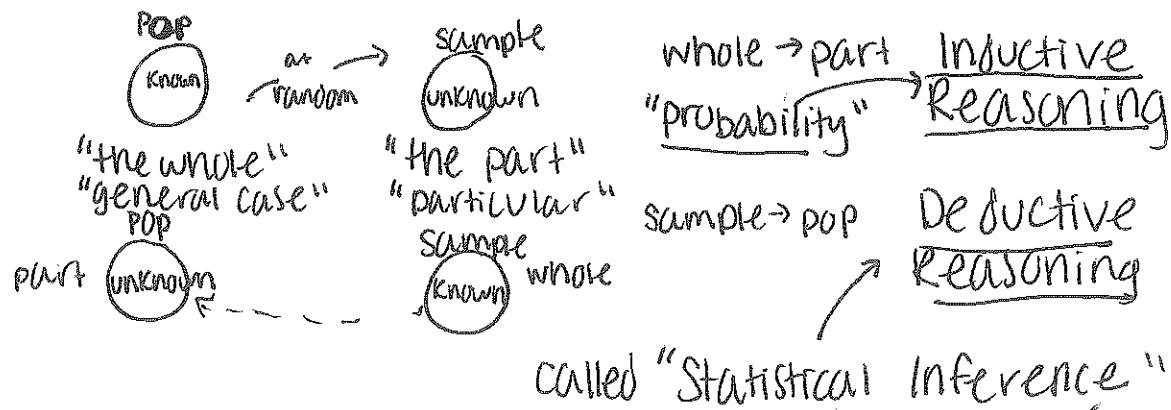


Crab case study continued:

Lets compare theoretical mean 24.3°C and experimental sample ~~25.0°C~~ 25.0°C .



Probability is easier than Statistical Inference.

Theory says $M = 24.3^{\circ}\text{C}$

Inferential Summary	
population \rightarrow	Unknown pop Summary of main interest $M = \text{pop mean equilibration } T^{\circ}\text{C}$
Sample \rightarrow	estimate of M (around) $\bar{y} = 25.0^{\circ}\text{C}$
image dataset	give or take \bar{y} as estimate of M $\hat{SE}(\bar{y}) = .27^{\circ}\text{C}$
	Confidence Interval for M $24.4^{\circ}\text{C}, 25.6^{\circ}\text{C}$

I think that M is around $\bar{y} = 25.0^{\circ}\text{C}$ give or take around $\hat{SE}(\bar{y}) = .27^{\circ}\text{C}$, and my confidence $24.4^{\circ}\text{C} + \pm 25.6^{\circ}\text{C}$

Now let's build

image data
all possible \bar{y} 's

$$\begin{bmatrix} \bar{y}_1 \\ \bar{y}_2 \\ \vdots \\ \bar{y}_M \end{bmatrix}$$

$$M \rightarrow \infty$$

hypothetical
IID

$$\begin{array}{c} \bar{y} = \\ \bar{y}_1 = \\ \vdots \\ \dots \end{array}$$

$$\begin{bmatrix} 25.0 \\ 24.8 \\ \vdots \end{bmatrix}$$

$$M \rightarrow \infty$$

$$= M$$

long run mean ($E\bar{y}$)

long run SD ($SE_{\bar{y}}$)

$$= \frac{\sigma}{\sqrt{n}} = \frac{s}{\sqrt{n}} = .27^\circ C$$

long run hist

$$= SE = .27$$



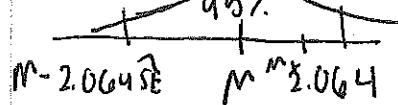
so let's fill in those blanks ↑

$E_{IID}(\bar{y}) = M$	MATH FACTS
$SE_{IID}(\bar{y}) = \frac{\sigma}{\sqrt{n}}$	CLT

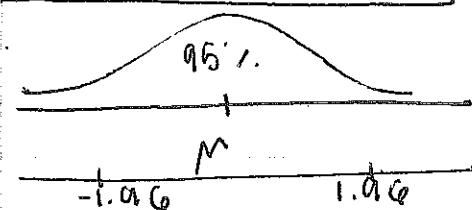
but let's get a number...

$$\text{Estimated } SE \text{ of } \bar{y} = SE_{IID}(\bar{y}) = \frac{s}{\sqrt{n}} = \frac{1.34^\circ C}{\sqrt{25}} = 0.27^\circ C$$

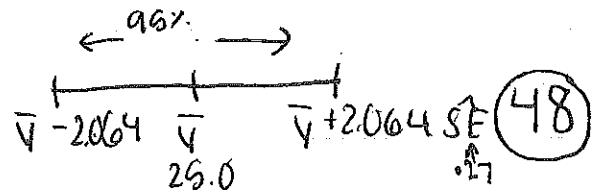
this is the "give or take" → $= 0.27^\circ C$
 $\hat{SE} = 0.27^\circ C$ normal long run hist. \bar{y}



use the t-table



$n-1$ degrees of freedom
as s as estimate of σ



$$\bar{Y} \pm (t_{n-1, .95}) \frac{s}{\sqrt{n}}$$

Neyman (1931)

interval of confidence
is a 95% confidence interval for M

$24.4^\circ C, 25.6^\circ C$

"The 95% interval"

Theory value for M was $24.3^\circ C$. Since $24.3^\circ C$ is not in the 95% interval for M , we conclude that the data set does not support theory at 95% confidence level.

When M_0 is outside 95% interval, people say the difference bw \bar{Y} and M_0 is statistically significant (statsig).

Significance:

- ① are $24.3^\circ C \neq 25.0^\circ C$
practically significantly different
 - if big enough to matter practically
 - more important, harder to judge
- ② are $24.3^\circ C \neq 25.0^\circ C$ stat sig different
(M_0 in 95% interval \rightarrow not statsig)

Highly statsig $\rightarrow 99\%$.

CI are about the process of building ~~more~~ intervals, not exact numbers.

Case Study: Maze w/ RATS

L = Left n = 12 rats

R = Right mean $\frac{10}{12} = 83\%$

Is it stat sig? Let's check (build interval)

(See L-157 for data sets)

Note: $\bar{p} = \hat{p} = \frac{10}{12}$

SD of 0/1 (pop w/ only 2 values):

$$(\text{larger value}) - (\text{smaller value}) \cdot \sqrt{\frac{\text{prob large} \cdot \text{prob small}}{p \cdot (1-p)}}$$

With pop 0/1 $\sigma = \sqrt{p(1-p)}$

MATH
FACT

Then build inferential summary table (see L-158)

$p = \text{pop \% that would turn left}$

imagine data $E\hat{p}$?
[] $SE \hat{p}$?

hist ?

} L-159

New Variables in this problem...

$$SE_{ID}(\hat{p}) = \sqrt{\frac{p(1-p)}{n}}$$

Estimated SE use $p = \hat{p} \rightarrow 11\%$

$\hat{p} = 83\% \quad \hat{SE}(\hat{p}) = 11\% \quad 95\% \text{ interval:}$

$$\hat{p} \pm 1.96 \hat{SE}(\hat{p})$$

$$\text{left} \quad 83 \quad 100\%, \quad 83 \pm 2(11\%) \rightarrow (61\%, 100\%)$$

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